

Chapter 3: Independence of Random Variables

Exercise 3.1 (Question 3.1.2) : Given n events $A_1, \dots, A_n \in \mathcal{A}$. Are $A_1, \dots, A_n$ independent events if only one of the following conditions holds?

- $\mathbb{P}(A_1 \cap \dots \cap A_n) = \mathbb{P}(A_1) \dots \mathbb{P}(A_n)$;
- for any pair $1 \leq i < j \leq n$, we have, $\mathbb{P}(A_i \cap A_j) = \mathbb{P}(A_i) \mathbb{P}(A_j)$.

Exercise 3.2 (Question 3.1.15) : Construct two random variables X and Y such that $\text{Cov}(X, Y) = 0$ without X and Y being independent. In Exercise 3.20, we will see a specific condition under which $\text{Cov}(X, Y) = 0$ implies the independence between X and Y .

Exercise 3.3 : Let X and Y be two independent standard normal random variables. Show that $\frac{X+Y}{\sqrt{2}}$ and $\frac{X-Y}{\sqrt{2}}$ are also independent with standard normal distribution.

Exercise 3.4 : Let X and Y be bounded random variables. Show that X and Y are independent if and only if

$$\forall k, \ell \in \mathbb{N}, \quad \mathbb{E}[X^k Y^\ell] = \mathbb{E}[X^k] \mathbb{E}[Y^\ell].$$

Exercise 3.5 : Let U be an exponential random variable with parameter 1 and V be a uniform random variable on $[0, 1]$. Assume that these two random variables are independent. Prove that $X = \sqrt{U} \cos(2\pi V)$ and $Y = \sqrt{U} \sin(2\pi V)$ are also independent.

Exercise 3.6 : Let X and Y be random variables with normal distribution with $\text{Cov}(X, Y) = 0$. Are they independent? If we add the condition “the random variable $aX + bY$ is also a normal distribution for all $a, b \in \mathbb{R}$ ”, please answer to the same question.

Exercise 3.7 (Question 3.1.17) : Let $X_1, \dots, X_n$ be real-valued random variables. The following properties are equivalent.

- $X_1, \dots, X_n$ are independent random variables.
- For any $a_1, \dots, a_n \in \mathbb{R}$, we have $\mathbb{P}(X_1 \leq a_1, \dots, X_n \leq a_n) = \prod_{i=1}^n \mathbb{P}(X_i \leq a_i)$.
- If $f_1, \dots, f_n$ are continuous and compactly supported (緊緻支撐) functions from $\mathbb{R}$ to $\mathbb{R}_{\geq 0}$, then

$$\mathbb{E} \left[\prod_{i=1}^n f_i(X_i) \right] = \prod_{i=1}^n \mathbb{E} [f_i(X_i)].$$

- The characteristic function of X writes,

$$\Phi_X(\xi_1, \dots, \xi_n) = \prod_{i=1}^n \Phi_{X_i}(\xi_i).$$

Exercise 3.8 (Question 3.1.19) : Let $\mathcal{B}_1, \dots, \mathcal{B}_n$ be independent σ -algebras and $n_0 = 0 < n_1 < \dots < n_p = n$. Then, the following σ -algebras are independent,

$$\begin{aligned}\mathcal{D}_1 &= \mathcal{B}_1 \vee \dots \vee \mathcal{B}_{n_1} \stackrel{(\text{def})}{=} \sigma(\mathcal{B}_1, \dots, \mathcal{B}_{n_1}), \\ \mathcal{D}_2 &= \mathcal{B}_{n_1+1} \vee \dots \vee \mathcal{B}_{n_2}, \\ &\vdots \\ \mathcal{D}_p &= \mathcal{B}_{n_{p-1}+1} \vee \dots \vee \mathcal{B}_{n_p}.\end{aligned}$$

Exercise 3.9 (Question 3.2.5) : Explain why it is necessary to assume that $(A_n)_{n \geq 1}$ is an independent sequence of events in (2) of Lemma 3.2.4.

Exercise 3.10 : Let $\alpha > 0$ and a sequence of independent random variables $(Z_n)_{n \geq 1}$ defined on the probability space $(\Omega, \mathcal{A}, \mathbb{P})$. Assume that for all $n \geq 1$, the distribution of Z_n is the following Bernoulli distribution,

$$\mathbb{P}(Z_n = 1) = \frac{1}{n^\alpha} \quad \text{and} \quad \mathbb{P}(Z_n = 0) = 1 - \frac{1}{n^\alpha},$$

Show that Z_n converges to 0 in L^1 but we have,

$$\text{a.s.,} \quad \limsup_{n \rightarrow \infty} Z_n = \begin{cases} 1 & \text{if } \alpha \leq 1, \\ 0 & \text{if } \alpha > 1. \end{cases}$$

Exercise 3.11 (Proposition 3.3.2) : If $(X_k)_{1 \leq k \leq n}$ is a sequence of i.i.d. random variables where each term follows the Poisson distribution of parameter λ , then $X_1 + \dots + X_n$ is a Poisson distribution of parameter $n\lambda$.

Exercise 3.12 (Proposition 3.3.3) : If $(X_k)_{1 \leq k \leq n}$ is a sequence of independent random variables such that for all $1 \leq k \leq n$, X_k has the Gaussian distribution of parameter $(0, \sigma_k^2)$, then $X_1 + \dots + X_n$ has the Gaussian distribution of parameter $(0, \sigma_1^2 + \dots + \sigma_n^2)$.

Exercise 3.13 : We use the notations from Exercise 2.15. Let $u > 0$. Assume that $X_1, \dots, X_n$ are i.i.d. Cauchy random variables with density function c_u , show that $\frac{1}{n}(X_1 + \dots + X_n)$ has the same density function c_u .

Exercise 3.14 (Gamma distribution) : Let $(X_n)_{n \geq 1}$ be an i.i.d. sequence of exponential random variables of parameter $\lambda > 0$. Given two parameters $k, \theta > 0$, if the random variable X has its values in $\mathbb{R}_{\geq 0}$ and the density function,

$$\gamma_{k,\theta}(x) = \frac{x^{k-1} e^{-x/\theta}}{\theta^k \Gamma(k)} \mathbf{1}_{x \geq 0}, \quad \forall x \in \mathbb{R}.$$

we call it the *Gamma distribution* (伽瑪分佈) of parameter (k, θ) , denoted $X \sim \Gamma(k, \theta)$.

(1) Given $k, \theta > 0$, compute the expectation and the variance of $\Gamma(k, \theta)$.

(2) Show that $X_1 + \dots + X_n \sim \Gamma(n, \lambda^{-1})$.

Exercise 3.15 : Let $X_1, \dots, X_n$ be i.i.d. standard Gaussian distributions. Show that the density function of $\chi_n^2 = X_1^2 + \dots + X_n^2$ is

$$\frac{x^{n/2-1} e^{-x/2}}{2^{n/2} \Gamma(n/2)} \mathbb{1}_{x>0}.$$

It is called the χ^2 distribution (χ^2 分佈) with n degrees of freedom. What is the relation of this distribution with the Gamma distribution (Exercise 3.14)?

Exercise 3.16 : Let $X_1, \dots, X_n$ be random variables defined on the probability space $(\Omega, \mathcal{A}, \mathbb{P})$. Assume that they are independent and have exponential distributions of parameters $c_1, \dots, c_n$. Let $Y = \min_{1 \leq k \leq n} X_k$. Prove the following statements.

- (1) Y is a random variable with exponential distribution of parameter $c_1 + \dots + c_n$.
- (2) There exists a $\mathcal{A}$ -measurable function $N : \Omega \longrightarrow \{1, \dots, n\}$ such that,

$$\mathbb{P}\text{-a.s.,} \quad X_N = Y \quad \text{and} \quad X_N < \min\{X_k : k \in \{1, \dots, n\} \setminus \{N\}\}.$$

- (3) The random variable N satisfies: for all $k \in \{1, \dots, n\}$, we have,

$$\mathbb{P}(N = k) = \frac{c_k}{c_1 + \dots + c_n}.$$

- (4) N and Y are independent random variables.

Exercise 3.17 (Poisson process) : Let $(\Omega, \mathcal{A}, \mathbb{P})$ be a probability space and $(X_n)_{n \geq 1}$ be a sequence of i.i.d. random variables with exponential distribution of parameter 1. Let $T_0 = 0$ and for all $n \geq 1$, let

$$T_n = X_1 + \dots + X_n.$$

For all $t \geq 0$, let

$$N_t = \max\{n \geq 0 : T_n \leq t\}.$$

- (1) Suppose $n \geq 1$, determine the distribution of the n -tuple $(T_1, \dots, T_n)$.
- (2) For any $t > 0$, determine the distribution of N_t .
- (3) Given $n \geq 1$ and $t > 0$, we define a new probability measure $\mathbb{Q}^{n,t}$ on Ω ,

$$\forall A \in \mathcal{A}, \quad \mathbb{Q}^{n,t}(A) = \frac{\mathbb{P}(A \cap \{N_t = n\})}{\mathbb{P}(N_t = n)}.$$

Determine the distribution of the n -tuple $(T_1, \dots, T_n)$ under the probability measure $\mathbb{Q}^{n,t}$.

Exercise 3.18 : We are given n independent random variables $X_1, \dots, X_n$ on the probability space $(\Omega, \mathcal{A}, \mathbb{P})$. Assume that they are all follow the uniform distribution on $[0, 1]$. Define $m = \min_{1 \leq i \leq n} X_i$ and $M = \max_{1 \leq i \leq n} X_i$. Determine the density functions of m and M , compute their expectations and variances.

Exercise 3.19 : Let $X = (X_1, \dots, X_d)$ be a d -dimensional real-valued random variable. We want to show that the following three conditions are equivalent. Moreover, when one of the conditions is satisfied, we say that X has a multivariate normal distribution (多元常態分佈).

- (a) There exist a d -dimensional real-valued random variable $Z = (Z_1, \dots, Z_d)$ such that the components are i.i.d. standard normal distributions, a square matrix A of size $d \times d$ and a vector $B \in \mathbb{R}^d$ such that $X \stackrel{(d)}{=} AZ + B$.
- (b) For any $\alpha \in \mathbb{R}^d$, the random variable $\alpha^T X$ also has a normal distribution.
- (c) There exist a semi-definite symmetric matrix Σ of size $d \times d$ and a vector $B \in \mathbb{R}^d$ such that the characteristic function of X writes,

$$\Phi_X(\xi) = \mathbb{E} [e^{i\xi \cdot X}] = \exp(i\xi^T B - \frac{1}{2}\xi^T \Sigma \xi).$$

Exercise 3.20 : Let (X, Y) be a pair of random variables with *bivariate normal distribution* (二元常態分佈). Prove that X and Y are independent if and only if $\text{Cov}(X, Y) = 0$.